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Midlife stroke risk and cognitive decline: A 10-year follow-up of the Whitehall II cohort study

Sara Kaffashian, MSc\textsuperscript{a,\ast}, Aline Dugravot, MSc\textsuperscript{a,b}, Eric J. Brunner, PhD\textsuperscript{c}, Séverine Sabia, PhD\textsuperscript{c}, Joël Ankri, MD\textsuperscript{b,d}, Mika Kivimäki, PhD\textsuperscript{c}, Archana Singh-Manoux, PhD\textsuperscript{a,b,c,d}

\textsuperscript{a}INSERM, U1018, Centre for Research in Epidemiology and Population Health, Villejuif, France
\textsuperscript{b}Université de Versailles, St Quentin, France
\textsuperscript{c}Department of Epidemiology and Public Health, University College London, UK
\textsuperscript{d}Centre de Gérontologie, Hôpital Ste Pétrine, Paris, France

Abstract

\textbf{Background:} Stroke is associated with an increased risk of dementia. However, it is unclear whether risk of stroke in those free of stroke, particularly in nonelderly populations, leads to differential rates of cognitive decline. Our aim was to assess whether risk of stroke in mid life is associated with cognitive decline over 10 years of follow-up.

\textbf{Methods:} We studied 4153 men and 1657 women (mean age, 55.6 years at baseline) from the Whitehall II study, a longitudinal British cohort study. We used the Framingham Stroke Risk Profile (FSRP), which incorporates age, sex, systolic blood pressure, diabetes mellitus, smoking, prior cardiovascular disease, atrial fibrillation, left ventricular hypertrophy, and use of antihypertensive medication. Cognitive tests included reasoning, memory, verbal fluency, and vocabulary assessed three times over 10 years. Longitudinal associations between FSRP and its components were tested using mixed-effects models, and rates of cognitive change over 10 years were estimated.

\textbf{Results:} Higher stroke risk was associated with faster decline in verbal fluency, vocabulary, and global cognition. For example, for global cognition there was a greater decline in the highest FSRP quartile ($-0.20 \pm 0.28$ to $-0.21$) compared with the lowest risk quartile ($-0.28$ to $-0.21$; $P = .03$). No association was observed for memory and reasoning. Of the individual components of FSRP, only diabetes mellitus was associated independently with faster cognitive decline ($\beta = -0.06$; 95\% confidence interval, $-0.01$ to 0.003; $P = .03$).

\textbf{Conclusion:} Elevated stroke risk at midlife is associated with accelerated cognitive decline over 10 years. Aggregation of risk factors may be especially important in this association.

\textsuperscript{\ast}Corresponding author. Tel.: +33 (0)1 77 74 74 06; Fax: +33 (0)1 77 74 74 03.
E-mail address: sara.kaffashian@inserm.fr

1. Introduction

Stroke increases the risk of dementia considerably [1–4]. In community-based studies, the prevalence of poststroke dementia in stroke survivors is about 30\% and the incidence of new onset dementia after stroke increases from 7\% after 1 year to 48\% after 25 years [5]. Cognitive impairment is three times more common in people who have had a stroke than in those who have not [6]. Even in the absence of stroke, individuals with a high risk of stroke have substantially higher cognitive deficits [7]. Individual risk factors such as obesity, hypertension, hypercholesterolemia, and smoking that predispose to stroke have been linked to adverse structural brain changes, cognitive impairment, and dementia [8–17]. Moreover, there is evidence that vascular risk factors predispose one to both vascular dementia and Alzheimer’s disease [5,18]. The clustering of risk factors may be particularly important and may increase the risk of cognitive impairment in an additive or synergistic manner, setting individuals on a trajectory of cognitive decline in advance of clinically detectable symptoms of cognitive impairment and dementia [16].

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Current practice guidelines on management and treatment of cardiovascular disease, both coronary heart disease and stroke, recommend use of multifactorial risk prediction models. These risk scores offer an effective evaluation of vascular risk particularly in younger populations, in which the prevalence of stroke is low, but risk of stroke may nevertheless be present as a result of the accumulation of low to moderate risk of multiple risk factors for stroke. Identification of associations of subclinical vascular risk with cognitive decline at the “brain at risk” stage is important and provides a great opportunity for prevention [19,20].

The Framingham risk scores are among the most widely validated and commonly used risk algorithms in clinical and research settings. However, their utility for predicting cognitive deficits in relation to vascular risk has remained relatively unexplored. The Framingham Stroke Risk Profile (FSRP) uses routinely measured risk factors to estimate 10-year risk of stroke. A number of studies have examined the relation of the FSRP in predicting cognitive deficits. The FSRP has been shown to be associated with incident cognitive impairment [21] and worse performance on various cognitive tests such as delayed verbal memory, verbal fluency [22], and abstract reasoning [7]. However, the few studies that have examined FSRP in relation to cognition are either limited by their cross-sectional design [7,22,23]; were conducted in a small, select population [24]; or had a short follow-up and used a single-item measure of cognitive status [21].

We sought to examine the association between stroke risk and longitudinal change in cognitive test scores using three repeated cognitive measures over a 10-year period in a large sample of middle-aged individuals.

2. Methods

2.1. Study population

The Whitehall II study was established in 1985 on 10,308 London-based office staff (6895 men and 3413 women). Details of the cohort and its follow-up have been described previously [25]. Briefly, all office staff aged 35 to 55 years in 20 civil service departments in London, UK, were invited to participate. In total, 73% of those invited agreed to participate in phase 1 (1985–1988), which consisted of a clinical examination and a self-administered questionnaire. Clinical examination included measures of blood pressure, anthropometry, biochemistry, neuroendocrine function, and subclinical markers of cardiovascular disease. Subsequent phases of data collection have alternated between a questionnaire-alone phase and a questionnaire accompanied by clinical examination. Cognitive testing was introduced to the full cohort during phase 5 (1997–1999) and was repeated during phase 7 (2002–2004) and phase 9 (2007–2009). All participants provided informed consent; the University College London ethics committee approved the study.

2.2. Assessment of Framingham Stroke Risk Profile

The FSRP is a clinical risk score that is used to calculate a sex-specific 10-year probability of stroke for individuals who are free of stroke at baseline. The original algorithm is based on prediction of 427 stroke events observed throughout a 10-year follow-up period for 2372 men and 3362 women in the Framingham Heart Study. The FSRP is based on the following risk factors: age, sex, systolic blood pressure, antihypertensive medication, diabetes, cigarette smoking status, history of cardiovascular disease, atrial fibrillation, and left ventricular hypertrophy as determined by electrocardiogram [26,27]. The FSRP has been shown to predict strongly the incidence of stroke in our cohort [28].

We drew risk score components from questionnaire and clinical examination data at phase 5 (1997–1999) in those free of stroke. Blood pressure was taken twice in the sitting position with a Hawksley random-zero sphygmomanometer after a 5-minute rest (Lynjay Services Ltd., Worthing, United Kingdom). The average of the two readings was used in the analysis. Antihypertensive medication included diuretics, beta blockers, angiotensin-converting enzyme inhibitors, and calcium channel blockers. Diabetes was defined by a fasting glucose level of ≥ 7.0 mmol/L, a 2-hour postload glucose level of ≥ 11.1 mmol/L, reported doctor-diagnosed diabetes, or use of diabetes medication [29]. Participants were categorized with respect to cigarette smoking status as current smokers or past/nonsmokers. History of cardiovascular disease was based on clinical examination at phases 1, 3, or 5 (via electrocardiograms and angiograms) and corroborated records obtained from general practitioners or hospitals. The diagnosis of atrial fibrillation and left ventricular hypertrophy was made using a standard 12-lead electrocardiogram and the Minnesota code classification system for electrocardiographic findings (atrial fibrillation, code 3-1 for High Amplitude R-waves; left ventricular hypertrophy, code B-3 for Arrhythmias). The stroke risk profile expressed as a predicted 10-year probability of incident stroke (measured as a percentage), was computed using beta coefficients based on the Cox proportional hazards regression model in the Framingham study [26,27].

2.3. Assessment of cognitive function

The cognitive test battery consisted of five standard tasks to assess performance in various cognitive domains. In our cohort, these tests were shown to be sensitive to detecting small changes in cognitive function over time in participants as young as 45 to 49 years of age [30].

The Alice Heim 4-I is composed of a series of 65 verbal and mathematical reasoning items of increasing difficulty [31]. It tests inductive reasoning, measuring the ability to identify patterns and infer principles and rules. The time allowed for this test was 10 minutes. Short-term verbal memory was assessed with a 20-word free recall test. Participants were presented a list of 20 one- or two-syllable words at
2-second intervals and were then asked to recall, in writing, as many of the words in any order, and they had 2 minutes to do so.

We used two measures of verbal fluency: phonemic and semantic. Phonemic fluency was assessed via “S” words and semantic fluency via “animal” words [32]. Participants were asked to recall in writing as many words beginning with “S” and as many animal names as they could. One minute was allowed for each test. Vocabulary was assessed using the Mill Hill Vocabulary test, used in its multiple-choice format, consisting of a list of 33 stimulus words ordered by increasing difficulty and six response choices [33]. In addition, a global cognitive score was created using all five of these tests by first standardizing raw scores on each test to z scores (mean, 0; standard deviation (SD), 1) using the baseline mean and SD value in the entire cohort for each test. Z scores were then averaged to yield the global cognitive score.

2.4. Covariates

Our analyses were adjusted for important demographic and health-related factors, including age, sex, ethnicity, education, depressive symptoms, physical activity, and alcohol use. Although age and sex are both components of the FSRP, we included them as covariates because of their established association with cognitive function. Age was centered at the mean value for the analytical sample. Ethnicity was categorized into (i) white and (ii) other ethnic groups. Education was measured as highest level of education achieved. Categories included (i) elementary or lower secondary, (ii) higher secondary (A’ levels), and (iii) first university degree or higher. We also assessed the effect of occupational position, classified as administrative, professional or executive, and clerical or supportive position. We also adjusted the analyses for depressive symptoms and health behaviors that in our cohort were shown to be associated with cognitive outcomes [34,35]. Depressive symptoms were assessed using the four-item depression subscale of the General Health Questionnaire and had two categories (0–3 and ≥4). Alcohol consumption was assessed using questions on the number of alcoholic drinks consumed during the past 7 days. This information was divided into “measures” of spirits, “glasses” of wine, and “pints” of beer, and was converted to number of units of alcohol, with each unit corresponding to 8 g ethanol. For example, a standard measure of spirits and a glass of wine are considered to contain 8 g alcohol, and a pint of beer, 16 g alcohol. Participants’ alcohol consumption was categorized as follows: none (0 unit/week), moderate (1–14 units/week for women and 1–21 units/week for men), and heavy (>14 units/week for women and >21 units/week for men). Physical activity level was determined from phase 5 questionnaires that included 20 items on frequency and duration of participation in different leisure-time physical activities (e.g., walking, general housework, cycling, sports) that were used to compute hours per week of each intensity level. Physical activity level was categorized as follows: high, more than 2.5 hours/week of moderate or more than 1 hour/week of vigorous physical activity; moderate, between 1 hour/week and 2.5 hours/week of moderate physical activity; and low, less than 1 hour/week of moderate and less than 1 hour/week of vigorous physical activity.

2.5. Statistical analysis

For ease of interpretation, we categorized stroke risk into quartiles, with the first quartile representing lowest risk and the fourth quartile representing highest stroke risk. To allow comparability of the five cognitive tests, they were standardized to z scores, as described previously.

We used linear mixed-effects models to estimate 10-year decline and associated 95% confidence intervals (CI) in each of the five measures of cognitive function. This method uses all available data over the 10 year follow-up, including those with one or two missing cognitive tests, and takes into account the intrapersonal correlation inherent in repeated measures. In our models, fixed effects included terms for time, the main effect term for each variable (stroke risk, age, sex, ethnicity, education, depressive symptoms, physical activity, and alcohol use), and interactions between time and each variable. The interaction between a given variable and time represents the effect of that variable on change in cognitive score over time. In addition to these main effects, two random effects were included—one for the intercept and one for the slope.

The stroke risk × time × sex interaction did not suggest sex differences in the rate of cognitive decline; thus, our analyses were not stratified by sex. Our final models adjusted for demographic factors included terms for FSRP quartile, time, age, sex, ethnicity, and education, and the interaction between time and each of FSRP quartile, age, sex, ethnicity, and education. An interaction term for FSRP quartile and age was also included. Additional models included terms for the demographics-adjusted models plus physical activity, alcohol use, and depressive symptoms, and their interaction with time. The low stroke risk quartile was used as the referent category to obtain P values for the difference in cognitive change in the remaining three quartiles. We conducted supplementary analyses to explore further each FSRP component as an independent risk factor for cognitive decline.

We also undertook sensitivity analyses to test the robustness of our main findings. We modeled the association between FSRP and cognitive decline, taking FSRP as a continuous variable, to represent a broader range of stroke risk. We carried out the same analyses as outlined earlier on log-transformed FSRP scores. We also conducted additional analyses to account for interim events (e.g., incident stroke) and missing cognitive function data in our analyses, and tested for the interaction with APOE ε4 carrier status (defined as presence or absence of ≥1 APOE ε4 allele) in a subset of the study population with these data (n = 4936).
Analyses were performed using Proc Mixed procedure in SAS software (version 9; SAS Institute Inc., Cary, NC, USA).

3. Results

Of 10,308 participants at the inception of the Whitehall II study (1985–1988), 7830 (75.9%) participated in phase 5 (1997–1999), the baseline of the current analysis. We included participants for whom stroke risk could be calculated (1997–1999), the baseline of the current analysis. We in-

clude participants for whom stroke risk could be calculated at phase 5 and who had at least one cognitive test measure during the follow-up (n = 6118). Our analysis is based on 5810 individuals (4153 men and 1657 women), after excluding participants with prevalent stroke (n = 48) or those missing data for covariates (n = 260); 73% of participants had complete data at all three phases and 20.2% at two phases. Restricting the analyses to participants with complete data at all three phases did not change our results significantly. Compared with individuals not included in these analyses, our sample included individuals who were younger (55.6 years vs. 57.5 years at phase 5, P < .001) and more educated (28.0% vs. 15% with a university degree, P < .001). Among participants who had at least one cognitive measure over three waves, those with data at all three waves were different from those with data at one or two waves; they had a lower FSRP (4.3% vs. 5.1%, P < .001), were younger (55.2 years vs. 56.7 years, P < .001), and were primarily men (72% vs. 67%, P < .001). There was also a higher proportion of individuals with a university education (29.8% vs. 24.9%, P < .01). Mean FSRP in our analytical sample was 4.5% (SD, 3.5%; range, 1%–78%). Detailed characteristics of the study population are presented in Table 1 (see Supplementary Table 1 for quartile-specific characteristics).

Mean cognitive test scores (z scores) at baseline (phase 5) by FSRP quartile are presented in Fig. 1. Persons in the highest stroke risk quartile had lower mean cognitive test scores compared with those in the lowest risk quartile, and there was a negative trend in cognitive test scores across the four stroke risk groups (P for trend = .001 on all tests except vocabulary, when P = .07).

Estimates of 10-year change in cognitive z scores as a function of stroke risk at baseline are presented in Table 2. In models adjusted for demographic factors, increased stroke risk was associated with faster cognitive decline for verbal fluency, vocabulary, and global cognition. For example, compared with a decline of −0.32 SD unit per 10 years in phonemic fluency test, for persons in the lowest stroke risk quartile, cognitive decline was −0.39 SD (95% CI, −0.46 to −0.32) and −0.41 SD (95% CI, −0.48 to −0.34) over 10 years for those in the third and fourth quartile, respectively. Similarly, in models adjusted for demographic and health-related factors, those in the fourth FSRP quartile showed faster decline in verbal fluency, vocabulary, and global cognition. Compared with the lowest stroke risk quartile, there was a 0.04-SD unit faster decline in global cognition in the fourth stroke risk quartile compared with the lowest risk quartile (P = .03). These differences were evident only when comparing the fourth quartile (FSRP ≥ 6%) with the lowest risk referent quartile. Adjusting associations for occupation, an indicator of socioeconomic status (instead of education), did not change the results considerably.

We investigated further individual FSRP components as independent risk factors for cognitive decline over 10 years, relating each of the FSRP components to global cognitive function. In models adjusted for demographics and health-related factors, only diabetes (β = −0.06; 95% CI, −0.01

Table 1

<table>
<thead>
<tr>
<th>Variable</th>
<th>Data</th>
</tr>
</thead>
<tbody>
<tr>
<td>Framingham Stroke Risk Profile (FSRP), mean (SD)</td>
<td>4.5 (3.5)</td>
</tr>
<tr>
<td>FSRP components</td>
<td></td>
</tr>
<tr>
<td>Age, y, mean (SD)</td>
<td>55.6 (6.0)</td>
</tr>
<tr>
<td>Men, n (%)</td>
<td>4153 (71.5)</td>
</tr>
<tr>
<td>Systolic blood pressure, mmHg, mean (SD)</td>
<td>123.0 (16.4)</td>
</tr>
<tr>
<td>Antihypertensive medications, n (%)</td>
<td>733 (12.6)</td>
</tr>
<tr>
<td>Diabetes, n (%)</td>
<td>225 (3.9)</td>
</tr>
<tr>
<td>Current smoker, n (%)</td>
<td>563 (9.7)</td>
</tr>
<tr>
<td>History of CVD, n (%)</td>
<td>331 (5.7)</td>
</tr>
<tr>
<td>Atrial fibrillation, n (%)</td>
<td>29 (0.5)</td>
</tr>
<tr>
<td>Left ventricular hypertrophy, n (%)</td>
<td>290 (5.0)</td>
</tr>
<tr>
<td><strong>Covariates</strong></td>
<td></td>
</tr>
<tr>
<td>Education, n (%)</td>
<td></td>
</tr>
<tr>
<td>Lower primary/secondary</td>
<td>2554 (43.9)</td>
</tr>
<tr>
<td>A levels</td>
<td>1522 (26.2)</td>
</tr>
<tr>
<td>University</td>
<td>1734 (29.8)</td>
</tr>
<tr>
<td>White ethnicity, n (%)</td>
<td>5370 (92.4)</td>
</tr>
<tr>
<td>Depressive symptoms, n (%)</td>
<td>716 (12.3)</td>
</tr>
<tr>
<td>Alcohol use, n (%)</td>
<td></td>
</tr>
<tr>
<td>None</td>
<td>871 (14.9)</td>
</tr>
<tr>
<td>Moderate</td>
<td>3522 (60.6)</td>
</tr>
<tr>
<td>Heavy</td>
<td>1417 (24.4)</td>
</tr>
<tr>
<td>Physical activity, n (%)</td>
<td></td>
</tr>
<tr>
<td>Low</td>
<td>1737 (29.9)</td>
</tr>
<tr>
<td>Moderate</td>
<td>975 (16.8)</td>
</tr>
<tr>
<td>High</td>
<td>3098 (53.3)</td>
</tr>
</tbody>
</table>

Abbreviations: CVD, cardiovascular disease; SD, standard deviation.
to −0.03; \( P = 0.03 \) was associated independently and inversely with decline in global cognition over 10 years (Table 3).

Last, we obtained similar results in analyses of stroke risk as a continuous variable. Here, too, verbal fluency, vocabulary, and global cognition were associated with cognitive decline over 10 years (see Supplementary Table 2). Other sensitivity analyses to test the robustness of our results indicated that accounting for interim stroke events ( \( n = 146 \) ) did not influence the findings greatly. Restricting the analyses to participants with complete cognitive data at all three phases ( \( n = 4258 \) ) did not alter the results qualitatively. In addition, the association between 10-year risk of stroke and cognitive decline was not modified by APOE ε4 status.

### 4. Discussion

In this large cohort of middle-aged men and women, we found higher risk of stroke to be associated with faster decline in multiple cognitive domains assessed using a battery of cognitive tests administered three times over 10 years. There was faster decline in phonemic and semantic fluency, vocabulary, and global cognition in the highest risk quartile compared with the referent lowest risk quartile. Our study, based on longitudinal data, provides a less biased estimate of the causal association between stroke risk and cognitive decline than cross-sectional data. The relationship between vascular risk factors and cognition is likely to be bidirectional; cross-sectional data on elderly subjects do not allow the direction of the association to be established. In this study, we addressed whether stroke risk in mid life, when stroke is rare, is associated with more rapid cognitive decline. Our findings add to the literature linking vascular risk to cognitive impairment by providing evidence for the association between stroke risk in mid life and long-term cognitive decline. Using the FSRP as an aggregate measure of risk factors, we found that mid to low risk levels on more than one component of the stroke risk score can lead to

### Table 2

Cognitive change as a function of stroke risk (FSRP)

<table>
<thead>
<tr>
<th>Cognitive tests</th>
<th>1 (mean, 2.4%)</th>
<th>2 (mean, 4.0%)</th>
<th>3 (mean, 5.0%)</th>
<th>4 (mean, 9.3%)</th>
</tr>
</thead>
<tbody>
<tr>
<td>Adjusted for demographics</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Reasoning</td>
<td>−0.31 (−0.34, −0.28)</td>
<td>−0.30 (−0.34, −0.26)</td>
<td>−0.33 (−0.38, −0.29)</td>
<td>−0.33 (−0.37, −0.29)</td>
</tr>
<tr>
<td>Memory</td>
<td>−0.23 (−0.29, −0.17)</td>
<td>−0.18 (−0.26, −0.10)</td>
<td>−0.23 (−0.32, −0.14)</td>
<td>−0.20 (−0.24, −0.08)</td>
</tr>
<tr>
<td>Phonemic fluency</td>
<td>−0.32 (−0.37, −0.27)</td>
<td>−0.34 (−0.40, −0.28)</td>
<td>−0.39 (−0.46, −0.32)</td>
<td>−0.41 (−0.48, −0.34)*</td>
</tr>
<tr>
<td>Semantic fluency</td>
<td>−0.22 (−0.27, −0.17)</td>
<td>−0.26 (−0.33, −0.22)</td>
<td>−0.26 (−0.33, −0.19)</td>
<td>−0.30 (−0.36, −0.24)*</td>
</tr>
<tr>
<td>Vocabulary</td>
<td>0.05 (0.02, 0.08)</td>
<td>0.06 (0.02, 0.09)</td>
<td>0.03 (0.004, 0.08)</td>
<td>−0.008 (−0.04, 0.03)*</td>
</tr>
<tr>
<td>Global cognition</td>
<td>−0.21 (−0.23, −0.18)</td>
<td>−0.21 (−0.23, −0.18)</td>
<td>−0.23 (−0.26, −0.21)</td>
<td>−0.24 (−0.27, −0.21)*</td>
</tr>
<tr>
<td>Adjusted for demographics and health related factors</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Reasoning</td>
<td>−0.33 (−0.36, −0.29)</td>
<td>−0.32 (−0.36, −0.28)</td>
<td>−0.35 (−0.40, −0.30)</td>
<td>−0.35 (−0.40, −0.30)</td>
</tr>
<tr>
<td>Memory</td>
<td>−0.25 (−0.32, −0.18)</td>
<td>−0.20 (−0.28, −0.12)</td>
<td>−0.25 (−0.34, −0.16)</td>
<td>−0.18 (−0.27, −0.099)</td>
</tr>
<tr>
<td>Phonemic fluency</td>
<td>−0.33 (−0.39, −0.27)</td>
<td>−0.34 (−0.41, −0.28)</td>
<td>−0.39 (−0.47, −0.32)</td>
<td>−0.42 (−0.48, −0.13)*</td>
</tr>
<tr>
<td>Semantic fluency</td>
<td>−0.21 (−0.27, −0.16)</td>
<td>−0.27 (−0.33, −0.20)</td>
<td>−0.25 (−0.32, −0.18)</td>
<td>−0.29 (−0.36, −0.22)*</td>
</tr>
<tr>
<td>Vocabulary</td>
<td>0.04 (0.01, 0.08)</td>
<td>0.05 (0.01, 0.09)</td>
<td>0.03 (0.01, 0.08)</td>
<td>−0.01 (−0.05, 0.03)*</td>
</tr>
<tr>
<td>Global cognition</td>
<td>−0.21 (−0.24, −0.19)</td>
<td>−0.16 (−0.24, −0.18)</td>
<td>−0.24 (−0.27, −0.20)</td>
<td>−0.25 (−0.28, −0.21)*</td>
</tr>
</tbody>
</table>

Abbreviations: CI, confidence interval; FSRP, Framingham Stroke Risk Profile.

NOTE. Estimates derived from linear mixed models using three assessments over 10 years. Models adjusted for demographics include age centered at the mean, sex, ethnicity, and education. Health-related factors include depressive symptoms, physical activity, and alcohol consumption.

* \( P < 0.05 \), difference in mean cognitive change compared with referent quartile 1.

\( P < 0.01 \), difference in mean cognitive change compared with referent quartile 1.
a threshold of risk that is detrimental to cognitive aging. We found that, even in our relatively low-risk middle-aged population, this led to a detectable decline in cognitive performance over 10 years. Although our data do not allow us to determine an exact cutoff for stroke risk that clearly confers a greater risk of cognitive decline, we observed a faster rate of cognitive decline only in the highest risk quartile (FSRP ≥ 6%; mean, 9.3%) compared with the referent lowest risk quartile. Thus, it appears that in our study sample a stroke risk ≥6% is associated with more rapid cognitive decline.

Among individual risk factors, only diabetes was associated independently with cognitive decline. Although this association has not been found in some studies [21], the observed relation between diabetes and cognitive decline in our study is supported by a number of studies reporting associations of diabetes with longitudinal changes on magnetic resonance imaging markers of vascular brain injury [17], incident cognitive impairment and cognitive decline [13,36,37], and incident dementia [38]. There is compelling evidence for a causal relationship between alterations in glycemic control and subsequent brain ischemic and atrophic changes [39]. Although the absence of an independent association between other risk factors for stroke and cognitive decline may be related to study population attributes (e.g., lower prevalence of risk factors in the current study), it can also substantiate the suggestion that the combined influence of clustering of these risk factors may be particularly important in this association. Indeed, stroke risk factors are correlated highly with each other. Overall, it appears that FSRP is a good predictor of cognitive impairment and cognitive decline. Our findings support further the notion that stroke risk factors begin to exert their influence on cognition early and that they may act in an additive manner in affecting cognitive decline. Furthermore, although in our study the differences in rate of cognitive decline between risk groups may seem small, they are nonetheless important from a pathogenetic stand point, especially considering that the mean age in our population at stroke risk assessment was only 55 years.

Although the neuropsychological tests used in various studies are not identical, it appears that stroke risk commonly assessed using the FSRP is associated with cognitive function in multiple cognitive domains [7,22–24]. We found that higher stroke risk was associated with decline in all cognitive domains except memory and reasoning. A 3-year follow-up study of 235 healthy older men showed an association between FSRP and decline only on verbal fluency; no associations between FSRP and decline in memory or visual–spatial function were reported [24]. In contrast, in a study of 23,752 stroke-free individuals monitored for an average of 4 years, a higher FSRP was found to be related to incident cognitive impairment in global cognitive function that focused on memory [21]. Several studies suggest that vascular disease does not always affect memory, and that vascular risk factors may have a particularly detrimental effect on frontally mediated cognitive functions such as verbal fluency [13,40,41]. Deficits in executive function have long been recognized as a salient feature of cognitive impairment of vascular origin. In relation to FSRP, at least three other cross-sectional studies report similar associations between stroke risk and cognitive function. Elias et al. [7] reported associations between higher FSRP and deficits in multiple domains, including abstract reasoning and visual–spatial memory, but not verbal memory, in 2175 participants of the Framingham Offspring Study. Seshadri et al. [23] found similar associations between FSRP and visual–spatial and executive function, but not with verbal memory. These results conflict with another study that found an association between FSRP and both immediate and delayed verbal memory, and semantic verbal fluency [22]. However, it is important to note that the comparison of our findings with these studies is limited because of important differences in study methods, particularly their cross-sectional design.

The mechanisms underlying the association between vascular risk factors and cognitive function remain to be elucidated fully. However, studies suggest that subclinical cerebrovascular disease may present an important link between major risk factors for stroke and cognitive function. These mechanisms include silent cerebral infarctions, brain atrophy, and white matter abnormalities [23,42–44]. Hypertension in mid life is associated with accelerated white matter hyperintensity volume (WMHV); diabetes and smoking are linked to a more rapid increase in temporal horn volume, which is a surrogate marker of accelerated hippocampal atrophy. Smoking has also been linked to a marked decrease in total brain volume [17].

The presence of silent brain infarcts on magnetic resonance imaging is associated with worse performance on neuropsychological tests and a steeper decline in global cognitive function, memory performance, and psychomotor speed [43].

The strengths of our study include its large sample size, three cognitive measures over 10 years to study longitudinal cognitive decline, and use of a representative battery of cognitive tests, permitting examination of different and distinct aspects of cognition. However, there are several limitations. First, because Whitehall II participants were office-based staff, they may not be representative of the British population, thus potentially limiting the generalizability of our findings. Second, it is possible that observed associations between stroke risk and cognitive decline are underestimations because our study sample included participants with a more favorable demographic and stroke risk profile than those excluded from the analysis and the general population.

In conclusion, the observed associations between an elevated risk of stroke and long-term cognitive decline in multiple cognitive domains have important implications. Our findings provide some evidence that an aggregation of vascular risk is especially important in this association, and suggest that vascular risk assessment may be better determined using a multifactorial risk score. That this association is
evident in a relatively young population indicates early neurological consequences of vascular risk factors leading to cognitive decline. Public health implications are in the domain of early prevention and treatment of stroke risk factors in middle age to reduce or forestall cognitive decline and dementia. Dementia is characterized by a long preclinical phase; individuals who develop dementia later may show subtle cognitive changes as early as two decades before diagnosis [45]. Currently, there are no specific treatments for cognitive impairment or dementia. Given the important vascular contribution to cognitive impairment, detection and treatment of risk factors particularly at mid life may be most effective in the prevention or progression of cognitive impairment. Indeed, recent projections suggest that as many as half of the cases of Alzheimer’s disease might be prevented by risk factor reduction and could potentially prevent up to 3 million cases of Alzheimer’s disease worldwide [46].

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