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Trajectories of the Framingham general cardiovascular risk profile in midlife and poor motor function later in life: The Whitehall II study

Alexis Elbaz a,b,c,⁎, Martin J. Shipley b, Hermann Nabi a,b,c, Eric J. Brunner b, Mika Kivimäki b, Archana Singh-Manoux a,b,c,d

a INSERM, Centre for Research in Epidemiology and Population Health, U1018, Social and Occupational Determinants of Health, F-94807, Villejuif, France
b Department of Epidemiology and Public Health, University College London, UK
Univ Versailles St-Quentin, UMRS 1018, F-94807, Villejuif, France
c Centre de Gérontologie, Hôpital Ste Périne, AP-HP, France

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Stroke
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A B S T R A C T

Background: Vascular risk factors are associated with increased risk of cognitive impairment and dementia, but their association with motor function, another key feature of aging, has received little research attention. We examined the association between trajectories of the Framingham general cardiovascular disease risk score (FRS) over midlife and motor function later in life.

Methods: A total of 5376 participants of the Whitehall II cohort study (29% women) who had up to four repeat measures of FRS between 1991–1993 (mean age = 48.6 years) and 2007–2009 (mean age = 65.4 years) and without history of stroke or coronary heart disease in 2007–2009 were included. Motor function was assessed in 2007–2009 through objective tests (walking speed, chair rises, balance, finger tapping, grip strength). We used age- and sex-adjusted linear mixed models.

Results: Participants with poorer performances for walking speed, chair rises, and balance in 2007–2009 had higher FRS concurrently and also in 1991–1993, on average 16 years earlier. These associations were robust to adjustment for cognition, socio-economic status, height, and BMI, and not explained by incident mobility limitation prior to motor assessment. No association was found with finger tapping and grip strength.

Conclusions: Cardiovascular risk early in midlife is associated with poor motor performances later in life. Vascular risk factors play an important and under-recognized role in motor function, independently of their impact on cognition, and suggest that better control of vascular risk factors in midlife may prevent physical impairment and disability in the elderly.

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1. Introduction

In addition to being strong predictors of cardio- and cerebrovascular disease, vascular risk factors have been associated with aging phenotypes, including worse cognitive function [1] and dementia [2]. Motor impairment is another key aspect of the aging process and poor motor function has been linked to adverse health outcomes, including disability [3] and death [4]. However, the association of vascular risk factors with motor function has received little research attention. To date, there is some evidence linking individual risk factors (hypertension [5], diabetes [6]) or markers of subclinical atherosclerosis [7–9] to poorer motor function, but the combined effect of vascular risk factors remains unknown.

For better prediction of cardiovascular disease (CVD) risk and a more complete assessment of vascular burden, several risk algorithms encompassing multiple risk factors have been developed [10]. Scores derived from these algorithms predict the risk of CVD, stroke, dementia [11], and cognitive deficit [12,13], but to our knowledge their association with motor function has not been examined. Here, we examine the association between trajectories of the Framingham general cardiovascular disease risk score (FRS) [14] during midlife, using four assessments over 16 years, and motor function at the end of the follow-up.

2. Methods

2.1. Participants

The Whitehall II study is a longitudinal study of 10,308 civil servants [15]. All civil servants aged 35–55 years in 20 London based departments were invited to participate.
(participation rate, 73%). The baseline examination took place over 1985–1988. Subse-
ty College London ethics committee approved the study.

2.2. Motor function

Motor function was assessed in 2007–2009 through measures of walking speed, chair
ries, balance, grip strength, and finger tapping; while the first three tests involve several
ystems and represent global measures, the last two are taken at the upper limbs and rep-
resent more specific measures (muscle strength, psychomotor speed). A practice session
was allowed for all. Correlations between tests were weak to moderate (supplementary
Table 1).

Walking speed was measured at usual pace over a marked 8-ft (2.44 m) course. The
starting position was standing at the start of the course. A trained nurse walked behind
the participant and stopped timing when the participant’s foot hit the floor after the end
of the course. Three tests were conducted; walking speed was computed as 2.44 m divid-
ed by the mean of three measures (in seconds).

Time to complete 5 chair rises: participants sat on an armless chair with feet resting on
the floor and arms folded across their chest. They stood up without using their arms and
sat down five times as quickly as possible. Time needed to complete the five chair rises
was recorded. Participants (n = 6) not able to stand up five times were excluded.

Balance was assessed through a series of tests of varying difficulty (full- and semi-
tandem stands, one-leg balance with eyes open or closed). For the present analyses, we
used means of the full-standing and one-leg balance test with eyes open. Particip-
ants were first asked to perform a full-tandem stand (10 s). If they passed this test, they
proceeded to perform a one-leg balance test (30 s). Participants who failed either
test were deemed to have failed the balance test.

Grip strength (in kilograms; dominant hand) was measured using a Smedley hand grip
dynamometer adjusted to suit participants’ hands with participants seated, their elbow on
the table, forearm pointing upwards, and palm of the hand facing up. Participants were
asked to squeeze the dynamometer as hard as possible for 2 s. These tests were performed
with a one minute rest between each. Readings were rounded up to the nearest whole
number; the mean of the tests was used.

Finger tapping test: the number of taps during 10 s was recorded using an electronic
device (WPS Electronic tapping test) [16] with participants holding their dominant hand
palm down, fingers extended, keeping their hand and arm stationary, and tapping on
the lever using their index.

2.3. Framingham general cardiovascular disease risk score

The FRS was developed as part of the Framingham Heart study to assess general CVD
risk and risk of individual events (coronary, cerebrovascular, peripheral artery disease,
heart failure) [14]. It includes measures of age, HDL- and total cholesterol, systolic blood
pressure, cigarette smoking, and diabetes, and provides an estimate of the 10-year risk of
CVD.

Risk score components were drawn from questionnaires and clinical examination data
HDL cholesterol (mg/dL) were measured from blood collected after either an 8 h
fast, seated without smoking, in the supine position after a 5 min test with the Hawkley cap-
tero sphygmomanometer. Treated hypertension was determined according to antihyper-
tensive medication use. Participants were categorized as current smokers, ex- or non-
smokers. Diabetes was defined by fasting glucose ≥ 7.0 mmol/L, 2 h post-load glucose
≥ 11.1 mmol/L, doctor diagnosed diabetes, or use of diabetes medication [17].

2.4. Covariates

Individuals with prevalent or incident stroke or coronary heart disease (non-fatal
myocardial infarction, definite angina) between 1991–1993 and 2007–2009 were exclud-
ed as these conditions are known to affect motor performances. Myocardial infarction was
diagnosed based on clinical examination data, electrocardiograms, and medical records
[18]. Angina was assessed based on reports of symptoms and nitrate medication, with cor-
roboration in medical records or abnormalities on a resting electrocardiogram, exercise
electrocardiogram, or coronary angiogram. Classification was carried out independently
by two trained coders, with adjudication by a third party in the event of disagreement.

Stroke was self-reported and included history of stroke or transient ischemic attack.

At all waves, mobility limitations were assessed using questions on the ability to climb
several flights of stairs or to walk more than 1 mile. Socioeconomic status (SES) was defined
based on the highest 3-level British civil service employment grade achieved (high, ad-
ministrative; intermediate, professional or executive; low, clerical or support). Weight
and height were measured and body mass index (BMI) was calculated as weight divided
by height squared (kg/m²). Cognitive status was assessed using the Alice Heim 4-I (AH-I4)
test [19], which includes 65 verbal and mathematical reasoning items assessing inductive
reasoning by measuring the ability to identify patterns and infer principles and rules;
higher scores correspond to better function. Participants had 10 min to do this test.

2.5. Statistical analysis

Descriptive analyses were carried out to examine participants’ characteristics at each
wave terms, association data, motor tests and FRS in 2007–2009. Correlations between
z-scores of motor tests were examined through age- and sex-adjusted partial Spearman
correlations.

The association between the FRS and motor tests was examined separately for each
test to establish whether motor function was associated with FRS concurrently and with
FRS trajectories over 16 years prior to motor testing. We defined age- and sex-specific
quartiles (supplementary Table 2) for all tests with continuous measures (walking
speed, grip strength, finger tapping, chair rises), given that motor performances decreased
with age (p < 0.001) and were higher in men than women (p < 0.001) [20]. For balance
(binary measure), models were age- and sex-adjusted.

We used linear mixed models that take into account correlations between repeated
measures on the same individual, with FRS as the dependent variable and quartiles of
motor tests as independent variables; this approach allows examining FRS trajectories
prior to the measurement of motor function (2007–2009) as well as their concurrent
association. FRS was logarithmically transformed due to its skewed distribution; results
were back-transformed for graphs. Models were implemented with a backward time
scale, so that 2007–2009 corresponds to the baseline (time = 0) and participants are
tracked back until 1991–1993, approximately 16 years earlier. Time was divided by 10,
so that regression coefficients represent change in FRS over 10 years. Inspection of the
data showed that FRS change over time was not linear; we therefore included a quadratic
term for time. Both the intercept and slope (time) were fitted as random effects. The main
effect represents the mean FRS difference in 2007–2009 between the reference quartile
(best performance) and other quartiles. The interaction term between quartiles of motor
tests and time allows examining whether the association between FRS and motor function
changed over time; non-linear differences in change were allowed by including interac-
tion terms between the quadratic time term and quartiles of motor tests.

Lower SES, weight, and height are strongly associated with motor performances and
FRS [21]. Poorer cognitive function is also associated with higher FRS [13] and worse
motor function [22,23]. We examined whether our findings were explained by confound-
ing (SES, weight, height) or mediated by cognition by including the following covariates
in models: known effects together with their significant (p < 0.05) interactions with time:
SES (high vs intermediate/low), quartiles of BMI and height, the measure of cognition
(AH-I4). These analyses used time invariant covariates defined in 2007–2009, and were
replicated using time-dependent covariates; they were adjusted for age in 2007–2009
and sex and their interactions with time (linear, squared).

In sensitivity analyses, we examined the influence of incident mobility limitations by
the assessment of motor function by excluding participants who reported mobility
limitations (limited to climb several flights of stairs or to walk more than 1 mile) at least

As FRS is higher in men than women [13], and men perform better than women on motor
tests [20], we examined whether sex modified cross-sectional associations between the
FRS and motor tests. We also investigated whether age modified their association.

Two-tailed p-values ≤ 0.05 were considered to be statistically significant. Statistical
analyses were performed using SAS 9.2 (SAS Institute, Cary, North Carolina, USA).

3. Results

Of 10,308 participants of Whitehall II at inception (1985–1988), 8104 participated in the clinical
tion; compared to those who participated in 2007–2009, those who did not were older (p < 0.001), more often women (p < 0.001), and had higher FRS (p < 0.001) in 1991–1993. We excluded 725 partici-
pants with a history of CHD/stroke, and 124 participants who did not have FRS or motor data. Our analyses
are based on 5376 participants. The chair rise test was missing for 7.2% of the participants due to more
stringent exclusion criteria than for other tests.

Participants’ characteristics at four waves are shown in Table 1. 3250 (60%) participants had four FRS measurements, 1225 (23%) three, 633 (12%) two, and 268 (5%) one. Mean (SD) FRS increased from 8.6
(6.3) in 1991–1993 to 16.9% (10.6) in 2007–2009 (on average 16.8 years later) with the prevalence of vascular risk factors, besides smoking, also rising over this period. Higher age, male sex, and vascular
risk factors were strongly associated with higher cardiovascular risk (Table 2); after adjustment for age and sex, higher BMI, smaller height, lower SES, and worse cognitive function remained associated with higher risk. Participants who developed mobility limitations between

Geometric FRS means from 1991–1993 to 2007–2009 according to quartiles of motor tests are presented in Table 3. Differences between the top and bottom quartiles were larger for walking speed, chair
rises, and balance than for grip strength and finger tapping, and these differences increased somewhat over time. The results of modeling FRS trajectories are presented in Table 4 and Fig. 1. There was good agreement between observed (Table 3) and predicted differences (Fig. 1) in FRS between top and bottom quartiles of motor tests. For walking speed, chair rises, and balance, participants with poor performances in 2007–2009 had higher FRS concurrently as well as in 1991–1993, on average 16 years earlier. The largest differences were observed for walking speed, followed by balance, and finally chair rises. After back-transformation of the log-FRS, differences in FRS across groups tended to increase over time and be larger in 2007–2009 than in 1991–1993 (Fig. 1). FRS was not associated with measures taken at upper limbs (grip strength, finger tapping) at any time point. Supplementary Tables 3–7 describe associations between participants’ characteristics, including FRS components, and motor tests. Associations between FRS and walking speed, chair rises, and balance were explained by a single covariate; note that the weaker association between low total cholesterol and slow walking speed disappeared (p = 0.19) after adjustment for HDL-cholesterol. Finger tapping and grip strength were not associated with most FRS components.

Adjustment for cognitive status in 2007–2009 had little influence (supplementary Fig. 1, panel a). After adjustment for SES, BMI, and height, the association between FRS and motor tests decreased but remained statistically significant, except for balance in 2007–2009 (supplementary Fig. 1, panel b); analyses using time-dependent covariates yielded similar conclusions (data not shown). After exclusion of

### Table 1
Clinical characteristics of Whitehall II participants.

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>Year</th>
<th>N = 4699</th>
<th>N = 3954</th>
<th>N = 4427</th>
<th>N = 5129</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>1991–1993</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Women, % (95% CI)</td>
<td>74 (72–76)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean age (SE)</td>
<td>62.2 (0.1)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>High grade, % (95% CI)</td>
<td>51 (46–57)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean BMI (SE), kg/m²</td>
<td>24.7 (0.1)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean height (SE), cm</td>
<td>168.7 (0.2)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean SBP (SE), mm Hg</td>
<td>112.4 (0.4)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean DBP (SE), mm Hg</td>
<td>72.7 (0.3)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Antihypertensive drugs, %</td>
<td>53.5 (47–53)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Mean total cholesterol (SE), mmol/L</td>
<td>1.93 (0.01)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Current smokers, % (95% CI)</td>
<td>0.7 (0.4–1.0)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Diabetes, % (95% CI)</td>
<td>1.3 (0.8–1.7)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>No mobility limitation, % (95% CI)</td>
<td>43.7 (0.3)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

BMI, body mass index; SBP, systolic blood pressure; DBP, diastolic blood pressure; FRS, Framingham cardiovascular disease.

### Table 2

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>Framingham general cardiovascular risk score</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>&lt;10%</td>
</tr>
<tr>
<td>Women, % (95% CI)</td>
<td>N = 1446</td>
</tr>
<tr>
<td>Mean age (SE)</td>
<td>74 (72–76)</td>
</tr>
<tr>
<td>High grade, % (95% CI)</td>
<td>62.2 (0.1)</td>
</tr>
<tr>
<td>Mean BMI (SE), kg/m²</td>
<td>51 (46–57)</td>
</tr>
<tr>
<td>Mean height (SE), cm</td>
<td>24.7 (0.1)</td>
</tr>
<tr>
<td>Mean SBP (SE), mm Hg</td>
<td>168.7 (0.2)</td>
</tr>
<tr>
<td>Mean DBP (SE), mm Hg</td>
<td>112.4 (0.4)</td>
</tr>
<tr>
<td>Antihypertensive drugs, %</td>
<td>53.5 (47–53)</td>
</tr>
<tr>
<td>Mean total cholesterol (SE), mmol/L</td>
<td>1.93 (0.01)</td>
</tr>
<tr>
<td>Current smokers, % (95% CI)</td>
<td>0.7 (0.4–1.0)</td>
</tr>
<tr>
<td>Diabetes, % (95% CI)</td>
<td>1.3 (0.8–1.7)</td>
</tr>
<tr>
<td>No mobility limitation, % (95% CI)</td>
<td>43.7 (0.3)</td>
</tr>
</tbody>
</table>

BMI, body mass index; SBP, systolic blood pressure; DBP, diastolic blood pressure.

* p-Values computed using linear regression or analysis of covariance adjusted for age and sex with the log-transformed FRS as the dependent variable and participants’ characteristics as explanatory variables. Age- and sex-adjusted means (SE) and percentages (95% CI) are presented.

**Defined as no self-reported difficulty to climb several flights of stairs and to walk more than 1 mile before the motor assessment according to questionnaires from 1991–1993, 1997–1999, and 2002–2004.**
participants who reported mobility limitations, associations remained statistically significant in 1991–1993 for walking speed and balance (supplementary Fig. 1, panel c). There were no sex- (p-values > 0.10) or age-related differences (p-values > 0.10) in cross-sectional associations between motor tests and FRS.

4. Discussion

Higher Framingham cardiovascular risk was associated with worse motor function assessed through measures of walking speed, chair rises, and balance. Based on repeated FRS assessments, higher risk was associated with worse motor function 16 years later. This association was robust to adjustments for BMI, height, SES, and cognitive function; it was not explained by mobility limitations before motor assessment. Our findings suggest that exposure to vascular risk factors in midlife is associated with poor motor function at older ages. The corollary, that low vascular risk in midlife may delay loss of motor function and extend independent living in the elderly, could become a key clinical and public health perspective across rapidly aging societies.

It is well recognized that cerebral multi-infarct states can lead to gait disorders, but the contribution of vascular risk factors to age-related changes in motor function has received little research attention. Our findings show that high vascular risk is associated with poorer motor function. This is in agreement with previous studies, mainly based on walking speed, that reported associations of worse motor function with individual vascular risk factors (hypertension [5,24], diabetes [6], homocysteine [25–27], low HDL-cholesterol [28]) or markers of vascular aging (increased common carotid artery intima–media thickness [7,8], arterial stiffness [9]). Most prior studies were cross-sectional or based on a single measure of vascular risk, while we assessed cardiovascular risk four times over many years, thus minimizing the risk of measurement error and reverse causation. To our knowledge, our study is the first investigation of FRS in relation to motor function. This score is widely used in clinical practice and has the advantage of integrating several risk factors.

Several mechanisms may account for the association between vascular risk factors and poor motor function. First, vascular risk factors are associated with an increased risk of cardio- and cerebrovascular events which lead to poor motor function. We excluded participants with overt vascular disease and our findings cannot be explained by such events. It is possible, however, that vascular health contributes to this association given the link between vascular outcomes such as peripheral artery disease [29] and heart failure [30] and physical functioning. Second, vascular risk factors are associated with an increased risk of white matter lesions, a marker of vascular brain injury assessed through brain MRI. Higher volumes of white matter lesions, particularly in the periventricular region, are associated with poor motor function, probably by disrupting brain circuits involved in motor control [31,32], and represent a ‘central’ vascular component. Third, vascular risk factors are associated with worse cognitive performances [1], which are themselves associated with motor function [22,23]. However, our results remained unchanged in analyses adjusted for a cognitive test associated with FRS in our study [13]: the association between cardiovascular risk and motor function was not explained by cognitive status.

There was some heterogeneity across motor tests. The association with FRS was strongest for walking speed and balance, less pronounced for chair rises, and absent for finger tapping and grip strength. These differences likely reflect the fact that walking speed, balance, and chair rises represent general measures of motor function, since they involve several systems and functions, including cardiovascular function, while finger tapping and grip strength pertain to more specific functions (psychomotor speed, muscle function). Another potential explanation is that vascular risk factors increase the risk of peripheral neuropathy in diabetic patients [33], which could explain the apparent sparing of upper limbs; however, analyses excluding diabetic patients yielded similar conclusions.

The absolute difference in cardiovascular risk between quartiles of motor function is modest. For instance, for walking speed, the absolute difference in the 2007–2009 FRS between the worse and best performers was 2.0%. However, the difference in walking speed between the top and bottom quartiles is large (~0.6–0.7 m/s). A meta-analysis estimated that 0.1 m/s higher walking speed was associated with a 12% reduced risk of mortality [4]; therefore, a 0.6 m/s difference corresponds to a roughly 50% reduced risk of death and is clinically relevant as it would significantly impact cardiovascular mortality [34,35].

Previous studies that examined a vascular contribution to motor performances were performed in elderly subjects. Selection biases
Table 4

<table>
<thead>
<tr>
<th>Parameter</th>
<th>Walking speed (N = 5336) (m/s) (time in seconds)</th>
<th>Five chair rises (N = 4958) (failed vs passed)</th>
<th>Balance test (N = 5268) (failed vs passed)</th>
<th>Finger tapping test (N = 5348) (number of taps per 10 s)</th>
<th>Grip strength (N = 5319) (kg)</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>Beta</td>
<td>95% CI</td>
<td>p</td>
<td>Beta</td>
<td>95% CI</td>
</tr>
<tr>
<td>Intercept</td>
<td>-2.022</td>
<td>-2.056; -1.987</td>
<td>&lt;10^-4</td>
<td>-2.007</td>
<td>-2.043; -1.971</td>
</tr>
<tr>
<td>MT-Q4^a</td>
<td>0.000</td>
<td>(Reference)</td>
<td></td>
<td>0.000</td>
<td>(Reference)</td>
</tr>
<tr>
<td>MT-Q3</td>
<td>0.027</td>
<td>-0.022; 0.076</td>
<td></td>
<td>0.056</td>
<td>0.006; 0.107</td>
</tr>
<tr>
<td>MT-Q2</td>
<td>0.074</td>
<td>0.025; 0.123</td>
<td></td>
<td>0.054</td>
<td>0.003; 0.105</td>
</tr>
<tr>
<td>MT-Q1</td>
<td>0.139</td>
<td>0.090; 0.188</td>
<td>&lt;10^-4</td>
<td>0.062</td>
<td>0.011; 0.113</td>
</tr>
<tr>
<td>Time</td>
<td>0.200</td>
<td>0.153; 0.247</td>
<td>&lt;10^-4</td>
<td>0.316</td>
<td>0.267; 0.365</td>
</tr>
<tr>
<td>Time^2</td>
<td>-0.143</td>
<td>-0.169; -0.111</td>
<td>&lt;10^-4</td>
<td>-0.077</td>
<td>-0.104; -0.050</td>
</tr>
<tr>
<td>Time × MT-Q4</td>
<td>0.000</td>
<td>(Reference)</td>
<td></td>
<td>0.000</td>
<td>(Reference)</td>
</tr>
<tr>
<td>Time × MT-Q2</td>
<td>0.081</td>
<td>0.014; 0.148</td>
<td></td>
<td>-0.054</td>
<td>-0.124; 0.015</td>
</tr>
<tr>
<td>Time × MT-Q1</td>
<td>0.066</td>
<td>-0.002; 0.133</td>
<td></td>
<td>-0.056</td>
<td>-0.126; 0.014</td>
</tr>
<tr>
<td>Time × MT-Q4</td>
<td>0.076</td>
<td>0.008; 0.144</td>
<td></td>
<td>-0.103</td>
<td>-0.172; -0.033</td>
</tr>
<tr>
<td>Time × MT-Q4</td>
<td>0.000</td>
<td>(Reference)</td>
<td></td>
<td>0.000</td>
<td>(Reference)</td>
</tr>
<tr>
<td>Time × MT-Q3</td>
<td>0.045</td>
<td>0.008; 0.081</td>
<td></td>
<td>-0.043</td>
<td>-0.081; -0.005</td>
</tr>
<tr>
<td>Time × MT-Q2</td>
<td>0.027</td>
<td>-0.009; 0.064</td>
<td></td>
<td>-0.038</td>
<td>-0.076; 0.000</td>
</tr>
<tr>
<td>Time × MT-Q1</td>
<td>0.039</td>
<td>0.001; 0.076</td>
<td>0.095</td>
<td>-0.058</td>
<td>-0.096; -0.020</td>
</tr>
</tbody>
</table>

Trajectories of the log-transformed FRS were modeled using linear mixed models with random effects for the intercept and time (linear, squared) and a backwards time scale. The main effects of the quartiles of the motor tests represent the difference in the log-transformed FRS between each quartile and the reference quartile in 2007–2009, and the interactions of the quartiles of the motor tests with time and time squared examine whether trajectories of the FRS are different in each quartile compared to the reference quartile. Regression coefficients (beta) were back transformed to present these findings graphically (Fig. 1).

a MT-Q1 to -Q4: age- and sex-specific quartiles of motor tests from worse (Q1) to better function (Q4, reference); for the balance test, Q4 corresponds to those who passed the test and Q3 to those who failed the test.

b p-Values for trend for the main effect of the quartiles of the motor tests and for their interactions with time (linear, squared) computed by including a four-level ordinal variable defined by the mean of the tests in each of the quartiles. We report a global test for the interactions of motor tests with time (linear, squared).

c The linear mixed model was adjusted for age (centered at 65 years; p < 10^-4) and sex (reference group, males; p < 10^-4), the interactions of age and sex with time and time squared (all p < 10^-4), the interaction between age and sex (p < 10^-4), and the three-way interaction between age, sex, and time (p = 0.002).
related to competing risks of death as well as reverse causation may complicate their interpretation. Our analyses, in contrast, are based on subjects who were on average ~49 years old at baseline followed for approximately 16 years; one important finding is that differences in FRS across groups defined by motor tests were already established many years before the motor assessment and well before the start of old age. However, the main limitation of our analyses is that motor function was measured once, and we were not able to study the association between the FRS and change in motor function. This limits causal inference as we were not able to formally assess temporality and to examine when differences in motor function appeared among participants; results of analyses with exclusion of persons who developed mobility limitations are reassuring to this respect. Our analyses are based on participants who participated in 2007–2009, which raises the possibility that selection biases due to survival or other causes may impact our findings. This is however unlikely given that the association between the FRS and motor tests was not modified by age; because persons who did not participate in 2007–2009 had higher FRS in 1991–1993 and are likely to be in worse health and have worse motor function [36], our estimates of the association between motor function and FRS are likely to be conservative. Finally, Whitehall II participants are office-based civil servants and not fully representative of the British population, which may limit the generalizability of our findings; this is likely, however, to mainly affect the distribution of risk rather than the strength of associations.

Our study has several strengths, including its large size and extended follow-up with at least three FRS measures available for most participants, thus allowing modeling trajectories over time. The availability of tests assessing different aspects of motor function is also an important feature.

In conclusion, our findings are in line with research on cognitive function and dementia showing that higher vascular risk in midlife is associated with worse outcomes later in life [2]. Our study is the first to show that participants with poor motor function have higher cardiovascular risk and that this association was already present approximately 16 years before the assessment of motor function. These findings suggest that vascular risk factors play an important and under-recognized role in motor function, independently of their impact on cognition. Thus, better control of vascular risk factors in midlife may prevent physical impairment and disability in the elderly.

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Appendix A. Supplementary data

Supplementary data to this article can be found online at http://dx.doi.org/10.1016/j.ijcard.2013.12.051.

References